

# Quick start for the sommer package

*Giovanny Covarrubias-Pazaran*

*2018-11-26*

The sommer package was developed to provide R users a powerful and reliable multivariate mixed model solver. The package is focused in problems of the type  $p > n$  (more effects to estimate than observations) and its core algorithm is coded in C++ using the Armadillo library. This package allows the user to fit mixed models with the advantage of specifying the variance-covariance structure for the random effects, and specify heterogeneous variances, and obtain other parameters such as BLUPs, BLUEs, residuals, fitted values, variances for fixed and random effects, etc.

The purpose of this quick start guide is to show the flexibility of the package under certain common scenarios:

B1) Background on mixed models

B2) Background on covariance structures

- 1) Univariate homogeneous variance models
- 2) Univariate heterogeneous variance models
- 3) Univariate unstructured variance models
- 4) Multivariate homogeneous variance models
- 5) Multivariate heterogeneous variance models
- 6) Multivariate unstructured variance models
- 7) Random regression models
- 8) GWAS models
- 9) Including special functions
  - the major `vs()` function for special variance models and its auxiliars:
    - `at()` specific levels structure
    - `ds()` diagonal structure
    - `us()` unstructured
    - `cs()` customized structure
    - `overlay()` overlayed models
    - `spl2D()` two dimensional spline models
- 10) The specification of constraints
- 11) Final remarks

## B1) Background on mixed models

The core of the package are the `mmer2` (formula-based) and `mmer` (matrix-based) functions which solve the mixed model equations. The functions are an interface to call the NR Direct-Inversion Newton-Raphson (Tunnicliffe 1989; Gilmour et al. 1995; Lee et al. 2016) or the EMMA efficient mixed model association algorithm (Kang et al. 2008). From version 2.0, sommer can handle multivariate models. Following Maier et al. (2015), the multivariate (and by extension the univariate) mixed model implemented has the form:

$$y_1 = X_1\beta_1 + Z_1u_1 + \epsilon_1$$

$$y_2 = X_2\beta_2 + Z_2u_2 + \epsilon_2$$

...

$$y_i = X_i\beta_i + Z_iu_i + \epsilon_i$$

where  $y_i$  is a vector of trait phenotypes,  $\beta_i$  is a vector of fixed effects,  $u_i$  is a vector of random effects for individuals and  $e_i$  are residuals for trait 'i' ( $i = 1, \dots, t$ ). The random effects ( $u_1 \dots u_i$  and  $e_i$ ) are assumed

to be normally distributed with mean zero.  $X$  and  $Z$  are incidence matrices for fixed and random effects respectively. The distribution of the multivariate response and the phenotypic variance covariance ( $V$ ) are:

$$Y = X\beta + ZU + \epsilon_i$$

$$Y \sim \text{MVN}(X\beta, V)$$

$$\mathbf{Y} = \begin{bmatrix} y_1 \\ y_2 \\ \dots \\ y_t \end{bmatrix}$$

$$\mathbf{X} = \begin{bmatrix} X_1 & 0 & 0 \\ \vdots & \ddots & \vdots \\ 0 & 0 & X_t \end{bmatrix}$$

$$\mathbf{V} = \begin{bmatrix} Z_1 K \sigma_{g_1}^2 Z_1' + H \sigma_{\epsilon_1}^2 & \dots & Z_1 K \sigma_{g_{1,t}} Z_t' + H \sigma_{\epsilon_{1,t}}^2 \\ \vdots & \ddots & \vdots \\ Z_1 K \sigma_{g_{1,t}} Z_t' + H \sigma_{\epsilon_{1,t}}^2 & \dots & Z_t K \sigma_{g_t}^2 Z_t' + H \sigma_{\epsilon_t}^2 \end{bmatrix}$$

where  $K$  is the relationship or covariance matrix for the  $k$ th random effect ( $u=1, \dots, k$ ), and  $H=I$  is an identity matrix or a partial identity matrix for the residual term. The terms  $\sigma_{g_i}^2$  and  $\sigma_{\epsilon_i}^2$  denote the genetic (or any of the  $k$ th random terms) and residual variance of trait ‘ $i$ ’, respectively and  $\sigma_{g_{ij}}$  and  $\sigma_{\epsilon_{ij}}$  the genetic (or any of the  $k$ th random terms) and residual covariance between traits ‘ $i$ ’ and ‘ $j$ ’ ( $i=1, \dots, t$ , and  $j=1, \dots, t$ ). The algorithm implemented optimizes the log likelihood:

$$\log L = 1/2 * \ln(|V|) + \ln(X'VX) + Y'PY$$

where  $||$  is the determinant of a matrix. And the REML estimates are updated using a Newton optimization algorithm of the form:

$$\theta^{k+1} = \theta^k + (H^k)^{-1} * \frac{dL}{d\sigma_i^2} | \theta^k$$

Where,  $\theta$  is the vector of variance components for random effects and covariance components among traits,  $H^{-1}$  is the inverse of the Hessian matrix of second derivatives for the  $k$ th cycle,  $\frac{dL}{d\sigma_i^2}$  is the vector of first derivatives of the likelihood with respect to the variance-covariance components. The Eigen decomposition of the relationship matrix proposed by Lee and Van Der Werf (2016) was included in the Newton-Raphson algorithm to improve time efficiency. Additionally, the popular pin function to estimate standard errors for linear combinations of variance components (i.e. heritabilities and genetic correlations) was added to the package as well.

The function `mmer` takes the  $Z$ s and  $K$ s for each random effect and construct the necessary structure inside and estimates the variance components by ML/REML using any of the 4 methods available in sommer. The `mmer2` function is enabled to work in a model-based fashion so user don't have to build the  $Z$ 's and  $K$  matrices. Please refer to the canonical papers listed in the Literature section to check how the algorithms work. We have tested widely the methods to make sure they provide the same solution when the likelihood behaves well but for complex problems they might lead to slightly different answers. If you have any concern please contact me at [cova\\_ruber@live.com.mx](mailto:cova_ruber@live.com.mx).

In the following section we will go in detail over several examples on how to use mixed models in univariate and multivariate case and their use in quantitative genetics.

## B2) Background on covariance structures

One of the major strenghts of linear mixed models is the flexibility to specify variance-covariance structures at all levels. In general, variance structures of mixed models can be seen as tensor (kronecker) products of multiple variance-covariance stuctures. For example, a multi-response model (i.e. 2 traits) where “g” individuals (i.e. 100 genotypes) are tested in “e” treatments (i.e. 3 environments), the variance-covariance for the random effect “individuals” can be seen as the following multiplicative model:

$$\mathbf{T} \otimes \mathbf{G} \otimes \mathbf{A}$$

where:

$$\mathbf{T} = \begin{bmatrix} \sigma_{g_{t1,t1}}^2 & \sigma_{g_{t1,t2}} \\ \sigma_{g_{t2,t1}} & \sigma_{g_{t2,t2}}^2 \end{bmatrix}$$

is the covariance structure for individuals among traits.

$$\mathbf{G} = \begin{bmatrix} \sigma_{g_{e1,e1}}^2 & \sigma_{g_{e1,e2}} & \sigma_{g_{e1,e3}} \\ \sigma_{g_{e2,e1}} & \sigma_{g_{e2,e2}}^2 & \sigma_{g_{e2,e3}} \\ \sigma_{g_{e3,e1}} & \sigma_{g_{e3,e2}} & \sigma_{g_{e3,e3}}^2 \end{bmatrix}$$

is the covariance structure for individuals among environments.

and  $\mathbf{A}$  is a square matrix representing the covariance among the levels of the individuals (any known relationship matrix).

The  $\mathbf{T}$  and  $\mathbf{G}$  covariance structures shown above are unknown matrices to be estimated whereas  $\mathbf{A}$  is known. The  $\mathbf{T}$  and  $\mathbf{G}$  matrices shown above are called as unstructured (US) covariance matrices, although this type is just one example from several covariance structures that the linear mixed models enable. For example, other popular covariance structures are:

Diagonal (DIAG) covariance structures

$$\mathbf{\Sigma} = \begin{bmatrix} \sigma_{g_{e1,e1}}^2 & 0 & 0 \\ \vdots & \ddots & \vdots \\ 0 & 0 & \sigma_{g_{ei,ei}}^2 \end{bmatrix}$$

Compound simmetry (CS) covariance structures

$$\mathbf{\Sigma} = \begin{bmatrix} \sigma_g^2 + \sigma_{ge}^2 & \sigma_g^2 & \sigma_g^2 & \sigma_g^2 \\ \sigma_g^2 & \sigma_g^2 + \sigma_{ge}^2 & \sigma_g^2 & \sigma_g^2 \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_g^2 & \sigma_g^2 & \sigma_g^2 & \sigma_g^2 + \sigma_{ge}^2 \end{bmatrix}$$

First order autoregressive (AR1) covariance structures

$$\mathbf{\Sigma} = \sigma^2 \begin{bmatrix} 1 & \rho & \rho^2 & \rho^3 \\ \rho & 1 & \rho & \rho^2 \\ \rho^2 & \rho & 1 & \rho \\ \rho^3 & \rho^2 & \rho & 1 \end{bmatrix}$$

or the already mentioned Unstructured (US) covariance structures

$$\Sigma = \begin{bmatrix} \sigma_{g_{e1,e1}}^2 & \sigma_{g_{e1,e2}} & \sigma_{g_{e1,e3}} \\ \vdots & \ddots & \vdots \\ \sigma_{g_{e3,e1}} & \sigma_{g_{e3,e2}} & \sigma_{g_{e3,e3}}^2 \end{bmatrix}$$

among others. Sommer has the capabilities to fit some of these covariance structures in the mixed model machinery.

## forming variance structures in sommer using the vs() function

The sommer function `vs()` allows to construct very structured variance models that are passed to the `mmer()` function it's one of the most important functions in the sommer package. Its specification is:

```
random=~vs(..., Gu, Gt, Gtc)
```

The idea is that the `vs()` function reflects the special variance structure that each random effect could have:

$T \otimes E \otimes \dots \otimes A$

where the `...` argument in the `vs()` function is used to specify the kronecker products from all matrices that form the variance for the random effect, where the auxiliary function `ds()`, `us()`, `cs()`, `at()`, can be used to define such structure. The idea is that a variance model for a random effect `x` (i.e. individuals) might require a more flexible model than just:

```
random=~x
```

For example, if individuals are tested in different time-points and environment, we can assume a different variance and covariance components among the individuals in the different environment-timepoint combinations. An example of variance structure of the type:

$T \otimes E \otimes S \otimes A$

would be specified in the `vs()` function as:

```
random=~vs(us(e),us(s),x, Gu=A, Gtc=T)
```

where the `e` would be a column vector in a data frame for the environments, `s` a vector in the dataframe for the time points, `x` is the vector in the dataframe for the identifier of individuals, `A` is a known square variance covariance matrix among individuals, and `T` is a square matrices with as many rows and columns as the number of traits.

### 1) Univariate homogeneous variance models

This type of models refer to single response models where a variable of interest (i.e. genotypes) needs to be analyzed as interacting with a 2nd random effect (i.e. environments), but you assume that across environments the genotypes have the same variance component. This is the so-called compound symmetry (CS) model.

```
library(sommer)
data(DT_example)
head(DT)
```

##	Name	Env	Loc	Year	Block	Yield	Weight
## 33	Manistee(MSL292-A)	CA.2013	CA	2013	CA.2013.1	4	-1.904711
## 65	CO02024-9W	CA.2013	CA	2013	CA.2013.1	5	-1.446958
## 66	Manistee(MSL292-A)	CA.2013	CA	2013	CA.2013.2	5	-1.516271
## 67	MSL007-B	CA.2011	CA	2011	CA.2011.2	5	-1.435510
## 68	MSR169-8Y	CA.2013	CA	2013	CA.2013.1	5	-1.469051

```
## 103          AC05153-1W CA.2013  CA 2013 CA.2013.1      6 -1.307167
```

```
ans1 <- mmer(Yield~Env,
             random= ~ Name + Env:Name,
             rcov= ~ units,
             data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)   restrained
##      1      -23.4747   21:31:8      0          0
##      2      -21.0636   21:31:8      0          0
##      3      -20.2304   21:31:8      0          0
##      4      -20.1465   21:31:8      0          0
##      5      -20.1454   21:31:8      0          0
##      6      -20.1454   21:31:8      0          0
```

```
summary(ans1)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value -20.14539 46.29075 55.95182      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## Name.Yield-Yield      3.682    1.6913  2.177    Positive
## Env:Name.Yield-Yield  5.173    1.4955  3.459    Positive
## units.Yield-Yield     4.366    0.6469  6.749    Positive
## =====
## Fixed effects:
##   Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept)    16.496    0.6855  24.065
## 2 Yield  EnvCA.2012    -5.777    0.7558  -7.643
## 3 Yield  EnvCA.2013    -6.380    0.7961  -8.015
## =====
## Groups and observations:
##           Yield
## Name         41
## Env:Name     123
## =====
## Use the '$' sign to access results and parameters
```

## 2) Univariate heterogeneous variance models

Very often in multi-environment trials, the assumption that the genetic variance or the residual variance is the same across locations may be too naive. Because of that, specifying a general genetic component and a location specific genetic variance is the way to go. This requires a CS+DIAG model (also called heterogeneous CS model).

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33  Manistee(MSL292-A) CA.2013  CA 2013 CA.2013.1      4 -1.904711
```

```
## 65          C002024-9W CA.2013  CA 2013  CA.2013.1      5 -1.446958
## 66  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.2      5 -1.516271
## 67          MSL007-B CA.2011   CA 2011  CA.2011.2      5 -1.435510
## 68          MSR169-8Y CA.2013  CA 2013  CA.2013.1      5 -1.469051
## 103         AC05153-1W CA.2013  CA 2013  CA.2013.1      6 -1.307167
```

```
ans2 <- mmer(Yield~Env,
             random= ~Name + vs(ds(Env),Name),
             rcov= ~ vs(ds(Env),units),
             data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)   restrained
##      1      -23.4747   21:31:8         0         0
##      2      -17.9787   21:31:9         1         0
##      3      -15.8441   21:31:9         1         0
##      4      -15.4545   21:31:9         1         0
##      5      -15.4314   21:31:9         1         0
##      6      -15.4299   21:31:9         1         0
##      7      -15.4298   21:31:9         1         0
```

```
summary(ans2)
```

```
## =====
##          Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##          logLik      AIC      BIC Method Converge
## Value -15.42992 36.85966 46.52072      NR      TRUE
## =====
## Variance-Covariance components:
##          VarComp VarCompSE Zratio Constraint
## Name.Yield-Yield      2.963   1.4969  1.979   Positive
## CA.2011:Name.Yield-Yield 10.145   4.5086  2.250   Positive
## CA.2012:Name.Yield-Yield  1.879   1.8704  1.005   Positive
## CA.2013:Name.Yield-Yield  6.629   2.5025  2.649   Positive
## CA.2011:units.Yield-Yield  4.942   1.5245  3.242   Positive
## CA.2012:units.Yield-Yield  5.724   1.3116  4.365   Positive
## CA.2013:units.Yield-Yield  2.560   0.6398  4.001   Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept)   16.508    0.8268  19.965
## 2 Yield  EnvCA.2012   -5.817    0.8575  -6.783
## 3 Yield  EnvCA.2013   -6.412    0.9356  -6.854
## =====
## Groups and observations:
##          Yield
## Name          41
## CA.2011:Name   41
## CA.2012:Name   41
## CA.2013:Name   41
## =====
## Use the '$' sign to access results and parameters
```

As you can see the special function `at` or `diag` can be used to indicate that there's a different variance for the genotypes in each environment. Same was done for the residual. The difference between `at` and `diag` is

that the `at` function can be used to specify the levels or specific environments where the variance is different.

### 3) Unstructured variance models

A more relaxed assumption than the CS+DIAG model is the unstructured model (US) which assumes that among the levels of certain factor (i.e. Environments) there's a covariance structure of a second random effect (i.e. Genotypes). This can be done in sommer using the `us(.)` function:

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.1      4 -1.904711
## 65           C002024-9W CA.2013  CA 2013  CA.2013.1      5 -1.446958
## 66  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.2      5 -1.516271
## 67           MSL007-B CA.2011  CA 2011  CA.2011.2      5 -1.435510
## 68           MSR169-8Y CA.2013  CA 2013  CA.2013.1      5 -1.469051
## 103          AC05153-1W CA.2013  CA 2013  CA.2013.1      6 -1.307167
```

```
ans3 <- mmer(Yield~Env,
             random=~ vs(us(Env),Name),
             rcov=~vs(us(Env),units),
             data=DT)
```

```
## iteration      LogLik      wall      cpu(sec)      restrained
##      1      -20.3368    21:31:9          0          0
##      2      -14.2124    21:31:9          0          0
##      3      -11.8334    21:31:9          0          0
##      4      -11.51     21:31:9          0          0
##      5      -11.5001    21:31:10         1          0
##      6      -11.4997    21:31:10         1          0
```

```
summary(ans3)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value -11.50013 28.99943 38.6605      NR      TRUE
## =====
## Variance-Covariance components:
##
##           VarComp VarCompSE      Zratio Constraint
## CA.2011:Name.Yield-Yield      15.6658 5.423e+00 2.889e+00 Positive
## CA.2012:CA.2011:Name.Yield-Yield      6.1107 2.486e+00 2.458e+00 Unconstr
## CA.2012:Name.Yield-Yield      4.5309 1.822e+00 2.487e+00 Positive
## CA.2013:CA.2011:Name.Yield-Yield      6.3844 3.067e+00 2.082e+00 Unconstr
## CA.2013:CA.2012:Name.Yield-Yield      0.3934 1.524e+00 2.581e-01 Unconstr
## CA.2013:Name.Yield-Yield      8.5972 2.484e+00 3.461e+00 Positive
## CA.2011:units.Yield-Yield      4.9698 1.532e+00 3.244e+00 Positive
## CA.2012:CA.2011:units.Yield-Yield      2.1067 1.437e-15 1.466e+15 Unconstr
## CA.2012:units.Yield-Yield      5.6722 1.300e+00 4.364e+00 Positive
## CA.2013:CA.2011:units.Yield-Yield      2.1067 0.000e+00      Inf Unconstr
## CA.2013:CA.2012:units.Yield-Yield      2.1067 0.000e+00      Inf Unconstr
## CA.2013:units.Yield-Yield      2.5571 6.392e-01 4.001e+00 Positive
```

```
## =====
## Fixed effects:
##   Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept) 16.331 0.8137 20.069
## 2 Yield EnvCA.2012 -5.696 0.7404 -7.693
## 3 Yield EnvCA.2013 -6.271 0.8191 -7.656
## =====
## Groups and observations:
##           Yield
## CA.2011:Name 41
## CA.2012:CA.2011:Name 82
## CA.2012:Name 41
## CA.2013:CA.2011:Name 82
## CA.2013:CA.2012:Name 82
## CA.2013:Name 41
## =====
## Use the '$' sign to access results and parameters
```

As can be seen the `us(Env)` indicates that the genotypes (Name) can have a covariance structure among environments (Env).

#### 4) Multivariate homogeneous variance models

Currently there's a great push for multi-response models. This is motivated by the correlation that certain variables hide and that could benefit in the prediction perspective. In sommer to specify multivariate models the response requires the use of the `cbind()` function in the response, and the `us(trait)`, `diag(trait)`, or `at(trait)` functions in the random part of the model.

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.1      4 -1.904711
## 65           CO02024-9W CA.2013  CA 2013  CA.2013.1      5 -1.446958
## 66  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.2      5 -1.516271
## 67           MSL007-B CA.2011  CA 2011  CA.2011.2      5 -1.435510
## 68           MSR169-8Y CA.2013  CA 2013  CA.2013.1      5 -1.469051
## 103          AC05153-1W CA.2013  CA 2013  CA.2013.1      6 -1.307167
```

```
DT$EnvName <- paste(DT$Env,DT$Name)
ans4 <- mmer(cbind(Yield, Weight) ~ Env,
             random= ~ vs(Name) + vs(EnvName),
             rcov= ~ vs(units),
             data=DT)
```

```
## iteration      LogLik      wall      cpu(sec)      restrained
##      1      72.7074 21:31:10      0      0
##      2     120.156 21:31:11      1      0
##      3     155.652 21:31:11      1      0
##      4     166.839 21:31:12      2      0
##      5     167.024 21:31:12      2      0
##      6     167.025 21:31:13      3      0
##      7     167.025 21:31:13      3      0
```

```
summary(ans4)
```



```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value 167.0252 -322.0505 -298.5695      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## u:Name.Yield-Yield      3.7090      1.68154 2.206 Positive
## u:Name.Yield-Weight      0.9071      0.37953 2.390 Unconstr
## u:Name.Weight-Weight      0.2243      0.08777 2.556 Positive
## u:EnvName.Yield-Yield      5.0922      1.47905 3.443 Positive
## u:EnvName.Yield-Weight      1.0269      0.30773 3.337 Unconstr
## u:EnvName.Weight-Weight      0.2101      0.06662 3.153 Positive
## u:units.Yield-Yield      4.3837      0.64949 6.750 Positive
## u:units.Yield-Weight      0.9077      0.14147 6.417 Unconstr
## u:units.Weight-Weight      0.2280      0.03378 6.751 Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept) 16.4093      0.6783 24.191
## 2 Weight (Intercept) 0.9806      0.1497 6.550
## 3 Yield EnvCA.2012 -5.6844      0.7474 -7.605
## 4 Weight EnvCA.2012 -1.1846      0.1593 -7.439
## 5 Yield EnvCA.2013 -6.2952      0.7850 -8.019
## 6 Weight EnvCA.2013 -1.3559      0.1681 -8.065
## =====
## Groups and observations:
##           Yield Weight
## u:Name      41      41
## u:EnvName    94      94
## =====
## Use the '$' sign to access results and parameters
```

You may notice that we have added the `us(trait)` behind the random effects. This is to indicate the structure that should be assumed in the multivariate model. The `diag(trait)` used behind a random effect (i.e. Name) indicates that for the traits modeled (Yield and Weight) there's no a covariance component and should not be estimated, whereas `us(trait)` assumes that for such random effect, there's a covariance component to be estimated (i.e. covariance between Yield and Weight for the random effect Name). Same applies for the residual part (`rcov`).

## 5) Multivariate heterogeneous variance models

This is just an extension of the univariate heterogeneous variance models but at the multivariate level. This would be a CS+DIAG multivariate model:

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33 Manistee(MSL292-A) CA.2013 CA 2013 CA.2013.1      4 -1.904711
## 65          C002024-9W CA.2013 CA 2013 CA.2013.1      5 -1.446958
## 66 Manistee(MSL292-A) CA.2013 CA 2013 CA.2013.2      5 -1.516271
```

```
## 67          MSL007-B CA.2011  CA 2011 CA.2011.2      5 -1.435510
## 68          MSR169-8Y CA.2013  CA 2013 CA.2013.1      5 -1.469051
## 103         AC05153-1W CA.2013  CA 2013 CA.2013.1      6 -1.307167
```

```
DT$EnvName <- paste(DT$Env,DT$Name)
ans5 <- mmer(cbind(Yield, Weight) ~ Env,
             random= ~ vs(Name) + vs(ds(Env),Name),
             rcov= ~ vs(ds(Env),units),
             data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)   restrained
##      1      72.7074  21:31:15         1           0
##      2     127.145  21:31:16         2           0
##      3     164.877  21:31:17         3           0
##      4     177.275  21:31:18         4           0
##      5     177.759  21:31:19         5           0
##      6     177.806  21:31:20         6           0
##      7     177.814  21:31:21         7           0
##      8     177.815  21:31:22         8           0
##      9     177.815  21:31:23         9           0
```

```
summary(ans5)
```

```
## =====
##          Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##          logLik      AIC      BIC Method Converge
## Value 177.8151 -343.6307 -320.1497      NR      TRUE
## =====
## Variance-Covariance components:
##
##          VarComp VarCompSE Zratio Constraint
## u:Name.Yield-Yield      3.31892  1.45282 2.2845  Positive
## u:Name.Yield-Weight      0.79393  0.32626 2.4335  Unconstr
## u:Name.Weight-Weight      0.19088  0.07504 2.5436  Positive
## CA.2011:Name.Yield-Yield      8.70864  4.01547 2.1688  Positive
## CA.2011:Name.Yield-Weight      1.77911  0.83930 2.1197  Unconstr
## CA.2011:Name.Weight-Weight      0.35964  0.17901 2.0090  Positive
## CA.2012:Name.Yield-Yield      2.57514  1.95105 1.3199  Positive
## CA.2012:Name.Yield-Weight      0.33336  0.39873 0.8360  Unconstr
## CA.2012:Name.Weight-Weight      0.03861  0.08602 0.4489  Positive
## CA.2013:Name.Yield-Yield      5.46595  2.16170 2.5285  Positive
## CA.2013:Name.Yield-Weight      1.34633  0.50444 2.6690  Unconstr
## CA.2013:Name.Weight-Weight      0.32882  0.12199 2.6955  Positive
## CA.2011:units.Yield-Yield      4.93854  1.52320 3.2422  Positive
## CA.2011:units.Yield-Weight      0.99447  0.32150 3.0932  Unconstr
## CA.2011:units.Weight-Weight      0.23982  0.07394 3.2433  Positive
## CA.2012:units.Yield-Yield      5.73847  1.31513 4.3634  Positive
## CA.2012:units.Yield-Weight      1.28000  0.30152 4.2452  Unconstr
## CA.2012:units.Weight-Weight      0.31804  0.07285 4.3655  Positive
## CA.2013:units.Yield-Yield      2.56126  0.63993 4.0024  Positive
## CA.2013:units.Yield-Weight      0.44569  0.12645 3.5246  Unconstr
## CA.2013:units.Weight-Weight      0.12232  0.03057 4.0009  Positive
## =====
## Fixed effects:
```

```
##      Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept) 16.4243 0.7891 20.814
## 2 Weight (Intercept) 0.9866 0.1683 5.863
## 3 Yield EnvCA.2012 -5.7339 0.8267 -6.936
## 4 Weight EnvCA.2012 -1.1998 0.1698 -7.066
## 5 Yield EnvCA.2013 -6.3129 0.8757 -7.209
## 6 Weight EnvCA.2013 -1.3621 0.1914 -7.115
## =====
## Groups and observations:
##      Yield Weight
## u:Name      41    41
## CA.2011:Name 41    41
## CA.2012:Name 41    41
## CA.2013:Name 41    41
## =====
## Use the '$' sign to access results and parameters
```

## 6) Multivariate unstructured variance models

This is just an extension of the univariate unstructured variance models but at the multivariate level. This would be a US multivariate model:

```
data(DT_example)
head(DT)
```

```
##      Name      Env Loc Year      Block Yield      Weight
## 33 Manistee(MSL292-A) CA.2013 CA 2013 CA.2013.1 4 -1.904711
## 65      CO02024-9W CA.2013 CA 2013 CA.2013.1 5 -1.446958
## 66 Manistee(MSL292-A) CA.2013 CA 2013 CA.2013.2 5 -1.516271
## 67      MSL007-B CA.2011 CA 2011 CA.2011.2 5 -1.435510
## 68      MSR169-8Y CA.2013 CA 2013 CA.2013.1 5 -1.469051
## 103     AC05153-1W CA.2013 CA 2013 CA.2013.1 6 -1.307167
```

```
DT$EnvName <- paste(DT$Env,DT$Name)
ans6 <- mmer(cbind(Yield, Weight) ~ Env,
             random= ~ vs(us(Env),Name),
             rcov= ~ vs(ds(Env),units),
             data=DT)
```

```
## iteration      LogLik      wall      cpu(sec)      restrained
##      1      82.6331 21:31:25      2      0
##      2     139.835 21:31:26      3      0
##      3     173.247 21:31:27      4      0
##      4     181.246 21:31:29      6      0
##      5     181.631 21:31:30      7      0
##      6     181.721 21:31:31      8      0
##      7     181.757 21:31:33     10      0
##      8     181.774 21:31:34     11      0
##      9     181.783 21:31:36     13      0
##     10     181.788 21:31:37     14      0
##     11     181.791 21:31:38     15      0
##     12     181.793 21:31:40     17      0
##     13     181.794 21:31:41     18      0
##     14     181.794 21:31:42     19      0
```

```
summary(ans6)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value 181.7937 -351.5889 -328.1079      NR      TRUE
## =====
## Variance-Covariance components:
##
##           VarComp VarCompSE Zratio Constraint
## CA.2011:Name.Yield-Yield      15.6464   5.35700  2.921   Positive
## CA.2011:Name.Yield-Weight       3.3589   1.14631  2.930   Unconstr
## CA.2011:Name.Weight-Weight       0.7182   0.24869  2.888   Positive
## CA.2012:CA.2011:Name.Yield-Yield   6.5290   2.48623  2.626   Unconstr
## CA.2012:CA.2011:Name.Yield-Weight   1.3505   0.52388  2.578   Unconstr
## CA.2012:CA.2011:Name.Weight-Weight   0.2842   0.11258  2.524   Unconstr
## CA.2012:Name.Yield-Yield       4.7893   1.86186  2.572   Positive
## CA.2012:Name.Yield-Weight       0.8640   0.38377  2.251   Unconstr
## CA.2012:Name.Weight-Weight       0.1693   0.08354  2.026   Positive
## CA.2013:CA.2011:Name.Yield-Yield   5.9932   2.93849  2.040   Unconstr
## CA.2013:CA.2011:Name.Yield-Weight   1.4232   0.64977  2.190   Unconstr
## CA.2013:CA.2011:Name.Weight-Weight   0.3379   0.14681  2.302   Unconstr
## CA.2013:CA.2012:Name.Yield-Yield   2.1001   1.44051  1.458   Unconstr
## CA.2013:CA.2012:Name.Yield-Weight   0.5243   0.32360  1.620   Unconstr
## CA.2013:CA.2012:Name.Weight-Weight   0.1343   0.07573  1.773   Unconstr
## CA.2013:Name.Yield-Yield       8.6267   2.47859  3.480   Positive
## CA.2013:Name.Yield-Weight       2.1050   0.58760  3.582   Unconstr
## CA.2013:Name.Weight-Weight       0.5126   0.14288  3.588   Positive
## CA.2011:units.Yield-Yield       4.9516   1.52694  3.243   Positive
## CA.2011:units.Yield-Weight       0.9993   0.32285  3.095   Unconstr
## CA.2011:units.Weight-Weight       0.2411   0.07432  3.244   Positive
## CA.2012:units.Yield-Yield       5.7791   1.32428  4.364   Positive
## CA.2012:units.Yield-Weight       1.2914   0.30409  4.247   Unconstr
## CA.2012:units.Weight-Weight       0.3212   0.07356  4.366   Positive
## CA.2013:units.Yield-Yield       2.5567   0.63883  4.002   Positive
## CA.2013:units.Yield-Weight       0.4452   0.12631  3.524   Unconstr
## CA.2013:units.Weight-Weight       0.1223   0.03056  4.001   Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept)  16.3342    0.8254  19.789
## 2 Weight (Intercept)   0.9677    0.1771   5.466
## 3 Yield EnvCA.2012  -5.6637    0.7449  -7.603
## 4 Weight EnvCA.2012  -1.1855    0.1604  -7.389
## 5 Yield EnvCA.2013  -6.2153    0.8340  -7.452
## 6 Weight EnvCA.2013  -1.3406    0.1806  -7.425
## =====
## Groups and observations:
##
##           Yield Weight
## CA.2011:Name      41    41
## CA.2012:CA.2011:Name  82    82
## CA.2012:Name      41    41
## CA.2013:CA.2011:Name  82    82
```

```
## CA.2013:CA.2012:Name      82      82
## CA.2013:Name              41      41
## =====
## Use the '$' sign to access results and parameters
```

Any number of random effects can be specified with different structures.

## 7) Random regression models

In order to fit random regression models the user can use the `leg()` function to fit Legendre polynomials. This can be combined with other special covariance structures such as `ds()`, `us()`, etc.

```
library(orthopolynom)
```

```
## Loading required package: polynom
```

```
data(DT_legendre)
head(DT)
```

```
##      SUBJECT X          Y Xf
## 1.1      1 1 -0.7432795  1
## 2.1      2 1 -0.6669945  1
## 3.1      3 1 -4.2802751  1
## 4.1      4 1  4.1092149  1
## 5.1      5 1 -3.0317213  1
## 6.1      6 1  1.3506577  1
```

```
mRR2<-mmer(Y~ 1 + Xf
           , random=~ vs(us(leg(X,1)),SUBJECT)
           , rcov=~vs(units)
           , data=DT)
```

```
## iteration   LogLik      wall   cpu(sec)   restrained
##      1      -166.081 21:31:43         0         0
##      2      -147.588 21:31:43         0         0
##      3      -137.701 21:31:43         0         0
##      4      -136.239 21:31:44         1         0
##      5      -136.222 21:31:44         1         0
##      6      -136.222 21:31:44         1         0
```

```
summary(mRR2)$varcomp
```

```
##              VarComp VarCompSE   Zratio Constraint
## leg0:SUBJECT.Y-Y    2.5783203 0.6717074 3.838458   Positive
## leg1:leg0:SUBJECT.Y-Y 0.4765587 0.2395195 1.989645   Unconstr
## leg1:SUBJECT.Y-Y    0.3497639 0.2183367 1.601947   Positive
## u:units.Y-Y         2.6911840 0.3824650 7.036420   Positive
```

Here, a numeric covariate X is used to explain the trajectory of the SUBJECT's and combined with an unstructured covariance matrix. The details can be found in the theory.

## 8) GWAS models

Although genome wide association studies can be conducted through a variety of approaches, the use of mixed models to find association between markers and phenotypes still one of the most popular approaches. Two of the most classical and popular approaches is to test marker by marker through mixed modeling (1 model by

marker) to obtain the marker effect and an statistic reflecting the level of association usually provided as the -log10 p-value. The second most popular approach is to assume that the genetic variance component is similar for all markers and therefore the variance components are only estimated once (1 model for all markers) and use the inverse of the phenotypic variance matrix ( $V^{-1}$ ) to test all markers in the generalized linear model  $b=(XV-X)-XV-y$ . This makes the GWAS much faster and efficient without major loses. Given the straight forward extension, sommer provides the GWAS function which can fit both type of approaches (be aware that these are 2 among many existant in the literature) in univariate and multivariate models, that way genetically correlated traits can be tested together to increase the power of detection.

Here we show a simple GWAS model for an univariate example.

```
data(DT_cpdata)
#### create the variance-covariance matrix
A <- A.mat(GT) # additive relationship matrix
#### look at the data and fit the model
head(DT,3)

##      id Row Col Year      color  Yield FruitAver Firmness Rowf Colf
## P003 P003   3   1 2014 0.10075269 154.67      41.93  588.917    3    1
## P004 P004   4   1 2014 0.13891940 186.77      58.79  640.031    4    1
## P005 P005   5   1 2014 0.08681502  80.21      48.16  671.523    5    1

head(MP,3)

##      Locus Position Chrom
## 1 scaffold_77830_839      0    1
## 2 scaffold_39187_895      0    1
## 3 scaffold_50439_2379      0    1

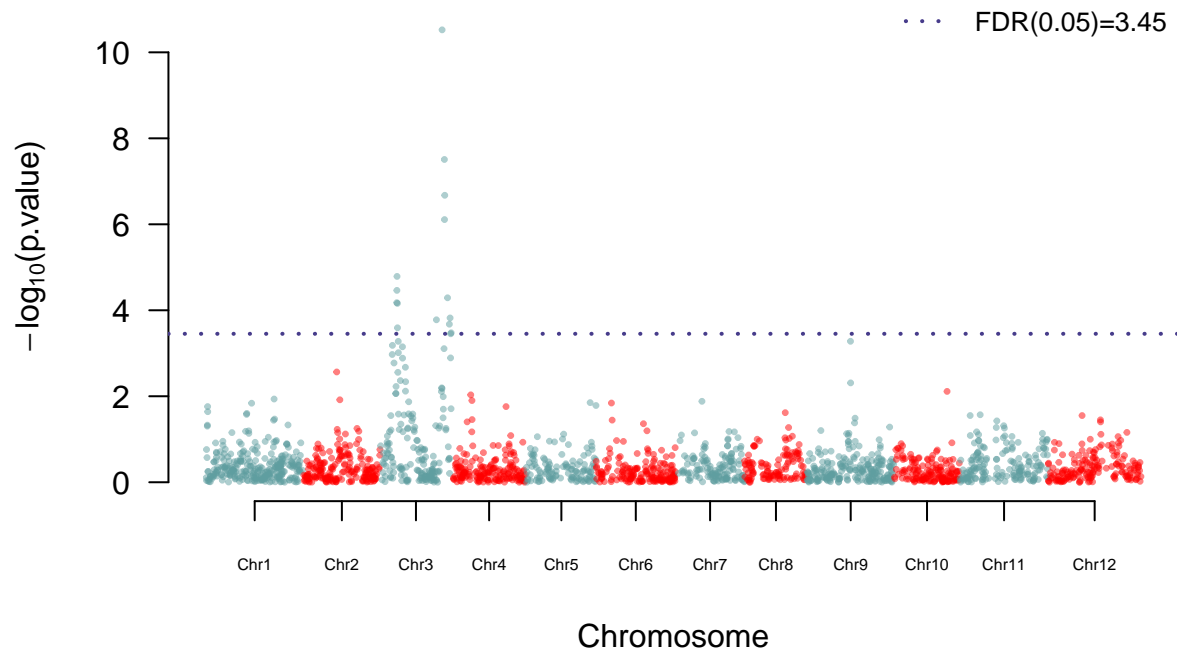
GT[1:3,1:4]

##      scaffold_50439_2381 scaffold_39344_153 uneak_3436043 uneak_2632033
## P003                  0                  0                0                1
## P004                  0                  0                0                1
## P005                  0                 -1                0                1

mix1 <- GWAS(color~1,
  random=~vs(id,Gu=A)
  + Rowf + Colf,
  rcov=~units,
  data=DT,
  M=GT, gTerm = "u:id")

## iteration   LogLik      wall    cpu(sec)   restrained
##      1      -147.849   21:31:45         0           0
##      2      -133.84   21:31:46         1           0
##      3      -117.12   21:31:46         1           0
##      4      -108.621  21:31:46         1           1
##      5      -108.14   21:31:47         2           1
##      6      -108.121  21:31:47         2           1
##      7      -108.12   21:31:47         2           1
## Performing GWAS evaluation

ms <- as.data.frame(t(mix1$scores))
ms$Locus <- rownames(ms)
MP2 <- merge(MP,ms,by="Locus",all.x = TRUE);
manhattan(MP2, pch=20,cex=.5, PVCN = "color score")
```



Be aware that the marker matrix  $M$  has to be imputed (no missing data allowed) and make sure that the number of rows in the  $M$  matrix is equivalent to the levels of the  $gTerm$  specified (i.e. if the  $gTerm$  is “id” and has 300 levels or in other words 300 individuals, then  $M$  has dimensions  $300 \times p$ , being  $p$  the number of markers).

## 9) Including special functions

Including special functions + the major `vs()` function for special variance models + `at()` specific levels structure + `ds()` diagonal structure + `us()` unstructured + `cs()` customized structure + `overlay()` overlaid models + `spl2D()` two dimensional spline models

In a mixed model framework there’s two types of covariance structures, the unknown and known. An example of a known covariance structure is the relationship matrix among individuals commonly present in plant and animal breeding programs. On the other hand, an example of an unknown covariance structure is in a multi-environment trial the covariance among genotypes in these environments, can be assumed diagonal, compound symmetry or unstructured but any needs to be estimated. In the following section we show how to specify unknown and known covariance structured for the random effects.

### the `vs()` function and its auxiliars `ds()`, `us()`, `at()` and `cs()`

The `vs()` function allows to fit different types of variance models (please take the time to read the documentation of this function). As explained in the introduction to covariance structures section in this document, the terms in the `vs()` function define the kronecker products that will be performed to define the variance and covariance components to be estimated. For example:

```
fixed=cbind(Y1,Y2,Y3)1 random=~vs(ds(Env),us(Time),Geno, Gu=A, Gtc=unsm(3)) rcov=~vs(ds(Env),us(Time),units)
```

defines a very complex model for the Geno random effect, where assumes that genotypes in different environments will be independent (diagonal structure using `ds()` function), but within each environment the different time points hold an unstructured variance-covariance structure (using the `us()` function), and at the same time a known covariance structure for Geno is specified in the `Gu` argument (here `A` is a square matrix provided by the user).

## the Gtc argument for constraints

At the same time all these is embebbbed in a multivariate model and the var-cov model is specified in the `Gtc` argument, here a full unstructured multivariate model is used by putting a 3x3 matrix in the `Gtc` argument with the following format:

$$\mathbf{Gtc} = \begin{bmatrix} 1 & 2 & 2 \\ 0 & 1 & 2 \\ 0 & 0 & 1 \end{bmatrix}$$

By default, sommer assumes an unstructured model if the `Gtc` argument is not provided. If the user wanted a DIAG model for the multivariate structure the argument would be `Gtc=diag(3)` which is again a 3x3 matrix but of a diagonal form:

$$\mathbf{Gtc} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

Remember that the numbers of the `Gtc` argument define the constraint applied in the model (1 positive, 2 unconstrained, 3 fixed)

Estimating a DIAG unknown covariance structure among genotypes in different environments (using the `ds()` function), same for residuals, and using a known covariance structure among genotypes (additive relationship matrix `A` applied in the `Gu` argument of the `vs` function).

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.1      4 -1.904711
## 65           C002024-9W CA.2013  CA 2013  CA.2013.1      5 -1.446958
## 66  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.2      5 -1.516271
## 67           MSL007-B CA.2011  CA 2011  CA.2011.2      5 -1.435510
## 68           MSR169-8Y CA.2013  CA 2013  CA.2013.1      5 -1.469051
## 103          AC05153-1W CA.2013  CA 2013  CA.2013.1      6 -1.307167
```

```
ans2 <- mmmer(Yield~Env,
              random= ~ vs(ds(Env),Name, Gu=A),
              rcov= ~ vs(ds(Env),units),
              data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)    restrained
##      1      -24.4279    21:32:3         0         0
##      2      -19.9461    21:32:3         0         0
##      3      -18.511    21:32:3         0         0
##      4      -18.3462    21:32:3         0         0
##      5      -18.3431    21:32:3         0         0
##      6      -18.343    21:32:3         0         0
```



```
summary(ans2)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value -18.34307 42.68598 52.34705      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## CA.2011:Name.Yield-Yield 17.214 6.1572 2.796 Positive
## CA.2012:Name.Yield-Yield 4.598 1.8373 2.503 Positive
## CA.2013:Name.Yield-Yield 8.790 2.5465 3.452 Positive
## CA.2011:units.Yield-Yield 4.954 1.5284 3.241 Positive
## CA.2012:units.Yield-Yield 5.662 1.2974 4.364 Positive
## CA.2013:units.Yield-Yield 2.557 0.6392 4.000 Positive
## =====
## Fixed effects:
##   Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept) 16.622 0.9485 17.525
## 2 Yield EnvCA.2012 -5.969 1.0447 -5.713
## 3 Yield EnvCA.2013 -6.659 1.0981 -6.064
## =====
## Groups and observations:
##           Yield
## CA.2011:Name 41
## CA.2012:Name 41
## CA.2013:Name 41
## =====
## Use the '$' sign to access results and parameters
```

and for multivariate models:

```
data(DT_example)
head(DT)
```

```
##           Name      Env Loc Year      Block Yield      Weight
## 33  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.1      4 -1.904711
## 65           CO02024-9W CA.2013  CA 2013  CA.2013.1      5 -1.446958
## 66  Manistee(MSL292-A) CA.2013  CA 2013  CA.2013.2      5 -1.516271
## 67           MSL007-B CA.2011  CA 2011  CA.2011.2      5 -1.435510
## 68           MSR169-8Y CA.2013  CA 2013  CA.2013.1      5 -1.469051
## 103          AC05153-1W CA.2013  CA 2013  CA.2013.1      6 -1.307167
```

```
ans2 <- mmer(cbind(Yield,Weight)~Env,
             random= ~ vs(ds(Env),Name, Gu=A, Gtc=unsm(2)),
             rcov= ~ vs(ds(Env),units, Gtc=diag(2)),
             data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)    restrained
##      1      73.0365  21:32:4      1          0
##      2      18.4838  21:32:5      2          0
##      3      68.5226  21:32:6      3          0
##      4      91.8798  21:32:6      3          0
##      5      92.4904  21:32:7      4          0
```

```
##      6      92.4963  21:32:8      5      0
##      7      92.4963  21:32:9      6      0
```

```
summary(ans2)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value 92.49628 -172.9927 -149.5116      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## CA.2011:Name.Yield-Yield  17.2100  6.12625  2.809  Positive
## CA.2011:Name.Yield-Weight   4.1997  1.30600  3.216  Unconstr
## CA.2011:Name.Weight-Weight   0.7979  0.28585  2.791  Positive
## CA.2012:Name.Yield-Yield   4.9115  1.87432  2.620  Positive
## CA.2012:Name.Yield-Weight   1.5623  0.36982  4.224  Unconstr
## CA.2012:Name.Weight-Weight   0.2031  0.08881  2.286  Positive
## CA.2013:Name.Yield-Yield   8.7891  2.53797  3.463  Positive
## CA.2013:Name.Yield-Weight   2.3723  0.60099  3.947  Unconstr
## CA.2013:Name.Weight-Weight   0.5259  0.14763  3.562  Positive
## CA.2011:units.Yield-Yield   4.8687  1.49431  3.258  Positive
## CA.2011:units.Weight-Weight  0.2363  0.07249  3.259  Positive
## CA.2012:units.Yield-Yield   5.4931  1.25622  4.373  Positive
## CA.2012:units.Weight-Weight  0.3031  0.06925  4.377  Positive
## CA.2013:units.Yield-Yield   2.5280  0.62975  4.014  Positive
## CA.2013:units.Weight-Weight  0.1209  0.03014  4.010  Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept)  16.623    0.9463  17.566
## 2 Weight (Intercept)   1.035    0.2044   5.065
## 3 Yield EnvCA.2012   -5.949    1.0454  -5.691
## 4 Weight EnvCA.2012  -1.251    0.2256  -5.547
## 5 Yield EnvCA.2013   -6.661    1.0958  -6.078
## 6 Weight EnvCA.2013  -1.445    0.2440  -5.923
## =====
## Groups and observations:
##           Yield Weight
## CA.2011:Name    41    41
## CA.2012:Name    41    41
## CA.2013:Name    41    41
## =====
## Use the '$' sign to access results and parameters
```

## customized random effects

One of the most powerful features of sommer is the ability to provide any customized matrix and estimate any random effect. For example:

```
data(DT_cpdata)
GT[1:4,1:4]
```

```
##      scaffold_50439_2381 scaffold_39344_153 uneak_3436043 uneak_2632033
## P003                0                0                0                1
## P004                0                0                0                1
## P005                0               -1                0                1
## P006               -1               -1               -1                0
```

#### look at the data and fit the model

```
mix1 <- mmer(Yield~1,
             random=~vs(list(GT)),
             rcov=~units,
             data=DT)
```

```
## iteration   LogLik      wall   cpu(sec)   restrained
##      1      -294.169  21:32:10         1         0
##      2      -240.607  21:32:10         1         0
##      3      -202.248  21:32:11         2         0
##      4      -180.448  21:32:11         2         0
##      5      -176.432  21:32:11         2         0
##      6      -176.212  21:32:11         2         0
##      7      -176.207  21:32:12         3         0
##      8      -176.207  21:32:12         3         0
```

the matrix GT is provided as a random effect by encapsulating the matrix in a list and provided in the `vs()` function.

## the `overlay()` function

Another very useful function is the `overlay` function, which allows to overlay matrices of different random effects and estimate a single variance component for the overlaid terms.

```
data("DT_halfdiallel")
head(DT)
```

```
##   rep geno male female    sugar
## 1   1  12    1      2 13.950509
## 2   2  12    1      2  9.756918
## 3   1  13    1      3 13.906355
## 4   2  13    1      3  9.119455
## 5   1  14    1      4  5.174483
## 6   2  14    1      4  8.452221
```

```
DT$femalef <- as.factor(DT$female)
DT$malef <- as.factor(DT$male)
DT$genof <- as.factor(DT$geno)
#### model using overlay
modh <- mmer(sugar~1,
             random=~vs(overlay(DT$femalef,DT$malef))
             + genof,
             data=DT)
```

```
## iteration   LogLik      wall   cpu(sec)   restrained
##      1      -7.04379  21:32:20         0         0
##      2      -6.09505  21:32:20         0         0
##      3      -5.71831  21:32:20         0         0
##      4      -5.67487  21:32:20         0         0
##      5      -5.67441  21:32:20         0         0
```

here the female and male random effects are overlayed becoming a single random effect that has the same variance component.

## the spl2D() function (using the 2-dimensional spline)

We will use the CPdata to show the use of 2-dimensional splines for accomodating spatial effects in field experiments. In early generation variety trials the availability of seed is low, which makes the use of unreplicated design a neccesity more than anything else. Experimental designs such as augmented designs and partially-replicated (p-rep) designs become every day more common this days.

In order to do a good job modeling the spatial trends happening in the field special covariance structures have been proposed to accomodate such spatial trends (i.e. autoregressive residuals; ar1). Unfortunately, some of these covariance structures make the modeling rather unstable. More recently other research groups have proposed the use of 2-dimensional splines to overcome such issues and have a more robust modeling of the spatial terms (Lee et al. 2013; Rodríguez-Álvarez et al. 2018).

In this example we assume an unreplicated population where row and range information is available which allows us to fit a 2 dimensional spline model.

```
data("DT_cpdata")
### mimic two fields
A <- A.mat(GT)
mix <- mmer(Yield~1,
            random=~vs(id, Gu=A) +
              vs(Rowf) +
              vs(Colf) +
              vs(spl2D(Row,Col)),
            rcov=~vs(units),
            data=DT)
```

##	iteration	LogLik	wall	cpu(sec)	restrained
##	1	-189.212	21:32:22	1	0
##	2	-168.339	21:32:22	1	0
##	3	-154.84	21:32:23	2	0
##	4	-151.445	21:32:23	2	0
##	5	-151.225	21:32:23	2	0
##	6	-151.203	21:32:24	3	0
##	7	-151.201	21:32:24	3	0
##	8	-151.201	21:32:25	4	0

```
summary(mix)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value -151.2012 304.4021 308.2937      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## u:id.Yield-Yield      783.3      319.2 2.4540 Positive
## u:Rowf.Yield-Yield      814.9      391.0 2.0840 Positive
## u:Colf.Yield-Yield      182.2      129.6 1.4056 Positive
## u:Row.Yield-Yield      513.4      694.4 0.7393 Positive
## u:units.Yield-Yield  2922.7      294.1 9.9365 Positive
```

```
## =====
## Fixed effects:
##   Trait      Effect Estimate Std.Error t.value
## 1 Yield (Intercept)    132.1     8.792   15.03
## =====
## Groups and observations:
##      Yield
## u:id      363
## u:Rowf     13
## u:Colf     36
## u:Row     168
## =====
## Use the '$' sign to access results and parameters
```

Notice that the job is done by the `spl2D()` function that takes the Row and Col information to fit a spatial kernel.

## 10) The specification of constraints

One of the major strengths of sommer is its extreme flexibility to specify variance-covariance structures in the multi-trait framework. Since sommer 3.7 this is easily achieved by the use of the `vs()` function and its argument `Gtc`. The idea behind how to specify the constraints has been explained in section 7) and here we will only show some examples.

Some useful function to create constrained matrices quickly are `unsm()` for unstructured, `uncm` for unconstrained, `fixm()` for fixed constraint, and `fcm()` for fixed effect constraints and its use is very easy:

```
unsm(4)
```

```
##      [,1] [,2] [,3] [,4]
## [1,]    1    2    2    2
## [2,]    2    1    2    2
## [3,]    2    2    1    2
## [4,]    2    2    2    1
```

can be used in `vs(x,Gtc=unsm(4))` to specify unstructured model for RE x

```
uncm(4)
```

```
##      [,1] [,2] [,3] [,4]
## [1,]    2    2    2    2
## [2,]    2    2    2    2
## [3,]    2    2    2    2
## [4,]    2    2    2    2
```

can be used in `vs(x,Gtc=uncm(4))` to specify unconstrained model for RE x

```
fixm(4)
```

```
##      [,1] [,2] [,3] [,4]
## [1,]    3    3    3    3
## [2,]    0    3    3    3
## [3,]    0    0    3    3
## [4,]    0    0    0    3
```

can be used in `vs(x,Gtc=fixm(4),Gt=mm)` to specify a fixed var-cov model for RE x and Gt needs to be provided

```
fcm(c(1,0,1,0))
```

```
##      [,1] [,2]
## [1,]    1    0
## [2,]    0    0
## [3,]    0    1
## [4,]    0    0
```

can be used in `vs(xf,Gtc=fcm(c(1,0,1,0)))` to specify that the fixed effect `xf` should be only estimated for traits 1 and 3

A matrix can combine the different constraints (0: not estimated, 1: positive, 2:unconstrained, 3:fixed) as desired.

First we show how to fit fixed effects for an specific trait. Here we assume two traits (Yield and Weight) and a fixed effect called “Env”, which we only want to fit for the trait number 2 and defaults for the random effects.

```
data(DT_example)
ansf <- mmer(cbind(Yield,Weight)~vs(Env,Gtc=fcm(c(0,1))),
             random= ~ vs(ds(Env),Name),
             rcov= ~ vs(ds(Env),units),
             data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)   restrained
##      1      33.8266  21:32:26      1          0
##      2     102.773  21:32:27      2          0
##      3     143.579  21:32:28      3          0
##      4     156.074  21:32:29      4          0
##      5     158.023  21:32:29      4          0
##      6     158.868  21:32:30      5          0
##      7     159.257  21:32:31      6          0
##      8     159.432  21:32:32      7          0
##      9     159.51  21:32:33      8          0
##     10     159.545  21:32:34      9          0
##     11     159.561  21:32:35     10          0
##     12     159.567  21:32:36     11          0
##     13     159.57  21:32:37     12          0
##     14     159.572  21:32:38     13          0
##     15     159.572  21:32:39     14          0
```

```
summary(ansf)
```

```
## =====
##           Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##           logLik      AIC      BIC Method Converge
## Value 159.5718 -311.1448 -295.4908      NR      TRUE
## =====
## Variance-Covariance components:
##           VarComp VarCompSE Zratio Constraint
## CA.2011:Name.Yield-Yield    51.6351  16.40155  3.148   Positive
## CA.2011:Name.Yield-Weight   11.0588   3.50866  3.152  Unconstr
## CA.2011:Name.Weight-Weight    2.3662   0.75361  3.140   Positive
## CA.2012:Name.Yield-Yield     4.5017   1.81385  2.482   Positive
## CA.2012:Name.Yield-Weight     0.8621   0.38456  2.242  Unconstr
```

```
## CA.2012:Name.Weight-Weight    0.1824    0.08619    2.117    Positive
## CA.2013:Name.Yield-Yield      9.1116    2.60811    3.494    Positive
## CA.2013:Name.Yield-Weight     2.2282    0.62058    3.590    Unconstr
## CA.2013:Name.Weight-Weight    0.5432    0.15137    3.588    Positive
## CA.2011:units.Yield-Yield     4.9660    1.53246    3.241    Positive
## CA.2011:units.Yield-Weight    1.0013    0.32371    3.093    Unconstr
## CA.2011:units.Weight-Weight   0.2413    0.07445    3.241    Positive
## CA.2012:units.Yield-Yield     5.6712    1.29969    4.364    Positive
## CA.2012:units.Yield-Weight    1.2616    0.29732    4.243    Unconstr
## CA.2012:units.Weight-Weight   0.3131    0.07174    4.365    Positive
## CA.2013:units.Yield-Yield     2.5508    0.63739    4.002    Positive
## CA.2013:units.Yield-Weight    0.4438    0.12600    3.522    Unconstr
## CA.2013:units.Weight-Weight   0.1220    0.03050    4.001    Positive
## =====
## Fixed effects:
##      Trait      Effect  Estimate Std.Error t.value
## 1  Yield (Intercept) 10.678099   0.33613 31.7678
## 2  Weight (Intercept) -0.233710   0.07736 -3.0209
## 3  Weight  EnvCA.2012  0.017250   0.04188  0.4118
## 4  Weight  EnvCA.2013 -0.008683   0.04116 -0.2110
## =====
## Groups and observations:
##      Yield Weight
## CA.2011:Name    41    41
## CA.2012:Name    41    41
## CA.2013:Name    41    41
## =====
## Use the '$' sign to access results and parameters
```

Now we specify an unstructured model for the random effect Name and the residuals and after a diagonal for both.

```
data(DT_example)
ans.uns <- mmer(cbind(Yield,Weight)~Env,
               random= ~ vs(Name,Gtc=unsm(2)),
               rcov= ~ vs(units,Gtc=unsm(2)),
               data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)    restrained
##      1      52.3292   21:32:39         0         0
##      2     107.273   21:32:40         1         0
##      3     142.651   21:32:40         1         0
##      4     154.389   21:32:41         2         0
##      5     154.654   21:32:41         2         0
##      6     154.655   21:32:41         2         0
##      7     154.655   21:32:42         3         0
```

```
summary(ans.uns)
```

```
## =====
##      Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##      logLik      AIC      BIC Method Converge
## Value 154.6554 -297.3108 -273.8298      NR      TRUE
## =====
```

```
## Variance-Covariance components:
##              VarComp VarCompSE Zratio Constraint
## u:Name.Yield-Yield    4.8593   1.52191   3.193   Positive
## u:Name.Yield-Weight    1.1432   0.34858   3.280   Unconstr
## u:Name.Weight-Weight    0.2737   0.08164   3.352   Positive
## u:units.Yield-Yield    8.1017   0.96030   8.437   Positive
## u:units.Yield-Weight    1.6523   0.20196   8.181   Unconstr
## u:units.Weight-Weight    0.3792   0.04497   8.432   Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
## 1  Yield (Intercept)  16.3396    0.5824  28.058
## 2  Weight (Intercept)   0.9641    0.1313   7.345
## 3  Yield  EnvCA.2012  -5.6429    0.5712  -9.878
## 4  Weight  EnvCA.2012 -1.1739    0.1245  -9.429
## 5  Yield  EnvCA.2013  -6.1768    0.6064 -10.186
## 6  Weight  EnvCA.2013  -1.3292    0.1327 -10.019
## =====
## Groups and observations:
##      Yield Weight
## u:Name    41    41
## =====
## Use the '$' sign to access results and parameters
```

```
ans.diag <- mmmer(cbind(Yield,Weight)~Env,
  random= ~ vs(Name,Gtc=diag(2)),
  rcov= ~ vs(units,Gtc=diag(2)),
  data=DT)
```

```
## iteration    LogLik      wall    cpu(sec)    restrained
##      1      52.3292  21:32:42      0          0
##      2     -73.3297  21:32:43      1          0
##      3     -65.0646  21:32:43      1          0
##      4     -63.8193  21:32:43      1          0
##      5     -63.815  21:32:44      2          0
##      6     -63.815  21:32:44      2          0
```

```
summary(ans.diag)
```

```
## =====
##      Multivariate Linear Mixed Model fit by REML
## ***** sommer 3.7 *****
## =====
##      logLik      AIC      BIC Method Converge
## Value -63.81504 139.6301 163.1111      NR      TRUE
## =====
## Variance-Covariance components:
##              VarComp VarCompSE Zratio Constraint
## u:Name.Yield-Yield    4.8559   1.52342   3.188   Positive
## u:Name.Weight-Weight    0.2733   0.08159   3.350   Positive
## u:units.Yield-Yield    8.1086   0.96146   8.434   Positive
## u:units.Weight-Weight    0.3793   0.04499   8.432   Positive
## =====
## Fixed effects:
##      Trait      Effect Estimate Std.Error t.value
```



```
## 1 Yield (Intercept) 16.3850 0.5849 28.012
## 2 Weight (Intercept) 0.9661 0.1313 7.359
## 3 Yield EnvCA.2012 -5.6880 0.5741 -9.908
## 4 Weight EnvCA.2012 -1.1756 0.1246 -9.437
## 5 Yield EnvCA.2013 -6.2183 0.6107 -10.182
## 6 Weight EnvCA.2013 -1.3304 0.1328 -10.021
## =====
## Groups and observations:
##      Yield Weight
## u:Name    41     41
## =====
## Use the '$' sign to access results and parameters
```

As a final example we will fit a multivariate model to deal with separate sexes which is a common problem in animal genetics.

```
# Generate some fake data:
# 100 males and 100 females
# Two traits are measured on each male, and two traits on each female
# 20 individuals per sex are measured for each of 5 different genotypes
set.seed(3434)
df <- data.frame(
  sex = rep(c("female", "male"), each = 100),
  female_trait_1 = c(rnorm(100), rep(NA, 100)),
  female_trait_2 = c(rnorm(100), rep(NA, 100)),
  male_trait_1 = c(rep(NA, 100), rnorm(100)),
  male_trait_2 = c(rep(NA, 100), rnorm(100)),
  genotype = rep(rep(1:5, each = 20), 2),
  individual = 1:200
)
df$genotype <- as.factor(df$genotype)
df$individual <- as.factor(df$individual)

mm <- adiag1(unsm(2),unsm(2));mm
```

```
##      [,1] [,2] [,3] [,4]
## [1,]    1    2    0    0
## [2,]    2    1    0    0
## [3,]    0    0    1    2
## [4,]    0    0    2    1
```

```
# mix <- mmer(cbind(female_trait_1,
#                   female_trait_2,
#                   male_trait_1,
#                   male_trait_2) ~ 1,
#                   random=~vs(genotype,Gtc=unsm(4)) + vs(individual,Gtc=mm),
#                   rcov=~vs(units), na.method.Y = "include",
#                   data=df)
# summary(mix)
```

I have silenced this code because data is not meaningful but this must show the way.

## 11) Final remarks

Keep in mind that sommer uses direct inversion (DI) algorithm which can be very slow for large datasets. The package is focused in problems of the type  $p > n$  (more random effect levels than observations) and models with dense covariance structures. For example, for experiment with dense covariance structures with low-replication (i.e. 2000 records from 1000 individuals replicated twice with a covariance structure of 1000x1000) sommer will be faster than MME-based software. Also for genomic problems with large number of random effect levels, i.e. 300 individuals ( $n$ ) with 100,000 genetic markers ( $p$ ). For highly replicated trials with small covariance structures or  $n > p$  (i.e. 2000 records from 200 individuals replicated 10 times with covariance structure of 200x200) asreml or other MME-based algorithms will be much faster and we recommend you to opt for those software.

## Literature

Covarrubias-Pazarán G. 2016. Genome assisted prediction of quantitative traits using the R package sommer. PLoS ONE 11(6):1-15.

Bernardo Rex. 2010. Breeding for quantitative traits in plants. Second edition. Stemma Press. 390 pp.

Gilmour et al. 1995. Average Information REML: An efficient algorithm for variance parameter estimation in linear mixed models. Biometrics 51(4):1440-1450.

Henderson C.R. 1975. Best Linear Unbiased Estimation and Prediction under a Selection Model. Biometrics vol. 31(2):423-447.

Kang et al. 2008. Efficient control of population structure in model organism association mapping. Genetics 178:1709-1723.

Lee, D.-J., Durban, M., and Eilers, P.H.C. (2013). Efficient two-dimensional smoothing with P-spline ANOVA mixed models and nested bases. Computational Statistics and Data Analysis, 61, 22 - 37.

Lee et al. 2015. MTG2: An efficient algorithm for multivariate linear mixed model analysis based on genomic information. Cold Spring Harbor. doi: <http://dx.doi.org/10.1101/027201>.

Maier et al. 2015. Joint analysis of psychiatric disorders increases accuracy of risk prediction for schizophrenia, bipolar disorder, and major depressive disorder. Am J Hum Genet; 96(2):283-294.

Rodriguez-Alvarez, Maria Xose, et al. Correcting for spatial heterogeneity in plant breeding experiments with P-splines. Spatial Statistics 23 (2018): 52-71.

Searle. 1993. Applying the EM algorithm to calculating ML and REML estimates of variance components. Paper invited for the 1993 American Statistical Association Meeting, San Francisco.

Yu et al. 2006. A unified mixed-model method for association mapping that accounts for multiple levels of relatedness. Genetics 38:203-208.

Abdollahi Arpanahi R, Morota G, Valente BD, Kranis A, Rosa GJM, Gianola D. 2015. Assessment of bagging GBLUP for whole genome prediction of broiler chicken traits. Journal of Animal Breeding and Genetics 132:218-228.

Tunncliffe W. 1989. On the use of marginal likelihood in time series model estimation. JRSS 51(1):15-27.